Lecture 10

Regularized Pricing and Risk Models Ivan Masyukov

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Plan for today

- Bonds
- Swaps
- Yield curve
- Regularized yield curve models
- Regularized volatility surface

Bonds

- A debt security
- Borrower issues bonds to obtain funds
- Investor purchases bond to earn return
- Typical bonds include fixed periodic coupon payments plus face value at maturity
- Zero coupon bonds only face value at maturity, no coupons
- There are perpetual bonds infinite regular coupon payments, but no face value, as the bonds never mature

Bond Cashflows

- Fixed rate bonds (periodic coupon payments and principal at maturity)
- Zero coupon bond
- Sum of future cashflows is not equal to bond price because future cashflowas are less valuable (time value of money)
- Discount factor

Bond Price

 Present price of the bond should be the sum of present values (PV) of future cashflows

$$P = \sum_{i=1}^{N} cF\Lambda_i + F\Lambda_N$$

- Where **P** fair bond price
 - **F** face value of bond
 - Λ_i discount factor for payment date *i*
 - c coupon rate
 - N number of coupon periods
- Need model for discounting Λ_i

Yield to Maturity

Use one parameter y – yield to maturity to compute all discount factors

$$\begin{split} \Lambda_{i} &= e^{-yt_{i}} \\ P &= e^{-yt_{1}}cF + e^{-yt_{2}}cF + \dots + e^{-yt_{N}}cF + e^{-yt_{N}}F \\ P &= \sum_{i=1}^{N} e^{-yt_{i}}C_{i} \end{split}$$

Where y - yield to maturity t - future time of payment, years $C_i - i-th cashflow$

- Continuous compounding case
- Assumed constant y for all ti

Bond Duration

• Sensitivity of bond price (In(P)) to bond yield

$$d = \frac{1}{P} \frac{\partial P}{\partial y}$$

$$d = -\frac{1}{P} \sum_{i=1}^{N} t_{i} e^{-yt_{i}} C_{i} = -\frac{\sum_{i=1}^{N} t_{i} e^{-yt_{i}} C_{i}}{\sum_{i=1}^{N} e^{-yt_{i}} C_{i}}$$
The d-bond duration

Where d - bond duration $C_i - i-th cashflow$

- Duration = "weighted time"
- Duration of zero coupon bond always equals to its maturity
- Duration of regular coupon bond is always less then its maturity
- As there is just one y for all payment dates, the duration is a sensitivity to "parallel" move

Bond Convexity

Second derivative of bond price to bond yield

$$c = \frac{\partial^2 P}{\partial y^2}$$
$$d = \sum_{i=1}^{N} t_i^2 e^{-yt_i} C_i$$

Where *c* – *bond convexity Ci* - *i-th cashflow*

- Duration is good measure for price changes for small variation in yield
- Second derivative needed for large changes in yields
- Convexity is always positive

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Fixed-vs-float swap analytics

Valuing fixed and float legs of the swap

$$PV _ fixed = \sum_{i} C\delta_{i}\Lambda_{i} = C\sum_{i} w_{i}$$
$$PV _ float = \sum_{i} r_{i}\delta_{i}\Lambda_{i} = \sum_{i} r_{i}w_{i}$$
$$PV _ fixed = PV _ float$$
$$C = \sum_{i} r_{i}w_{i} / \sum_{i} w_{i}$$

Where **C** – Swap rate (fixed leg coupon)

- Λ_i discount factor for payment date *i*
- δ_i day count fraction
- r_i forward rate (floating rate of future payment)
- Swap rate is weighted sum of forward rates (assumed same frequency of payments of fixed and floating legs)
- Swap can be hedged with bond

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Constructing Yield Curve

- Select input instruments
- Choose interpolation
 - Interpolation space (daily forward rates, zero rates, etc.)
 - Spline (piecewise-constant, linear, tension spline, etc.)
 - Knot points and model parameters
- Calibrate = solve for spline parameters such that input instruments are re-priced at par

Yield Curve Graph





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Bond Spread to Yield Curve

- We have curve now. So we can use can compute more accurate discount factors Λ_i , rather than relying on "flat" curve with same **y** for all cashflow dates
- Need extra parameter bond spread **s** to match with bond price

$$P = \sum_{i=1}^{N} e^{-st_i} \Lambda_i C_i$$

- Where Λ_i discount factor for payment date *i* computed from the curve s bond spread t_i future time of payment, years Ci *i*-th cashflow
 - If model is available for typical movements of the curve embedded in Λ_i we can build more effective risk model for bond, rather than using single "parallel" shift mode (bond duration)

Shifting 9Y swap by 1 basis point





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Portfolio Risk and Cost of Hedging

• Portfolio risk and Bid-Offer charge per bucket

Instrument	Quote	Raw Risk	B/O charge bp	Charge
IRS=1Y	0.33	(200,000)	0.10	20,000
IRS=2Y	0.39	1,330,000	0.10	133,000
IRS=3Y	0.49	(200,000)	0.25	50,000
IRS=4Y	0.64	1,200,000	0.25	300,000
IRS=5Y	0.86	(722,450)	0.10	72,245
IRS=6Y	1.09	(35,255)	0.25	8,814
IRS=7Y	1.29	(537,430)	0.25	134,358
IRS=8Y	1.48	(3,850,000)	0.25	962,500
IRS=9Y	1.64	1,580,000	0.25	395,000
IRS=10Y	1.79	288,751	0.10	28,875
IRS=12Y	2.04	(401,350)	0.25	100,338
IRS=15Y	2.29	50,000	0.25	12,500
IRS=20Y	2.50	4,000,000	0.25	1,000,000
IRS=25Y	2.60	(1,000,000)	0.25	250,000
IRS=30Y	2.67	(1,500,000)	0.10	150,000
TOTAL		2,266		3,617,629

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$$\mathbf{x} = \arg\min\{||\mathbf{F}^{\mathrm{T}}(\mathbf{r} + \mathbf{H}\mathbf{x})||^{2}\}$$

- r portfolio risk
- H hedging portfolio risks
- **x** weights of hedging instruments
- F market scenarios (factors)

Principal Component Analysis (PCA)

$\mathbf{D} = \mathbf{U}\mathbf{S}\,\mathbf{P}^{\mathrm{T}}$

- Use SVD to decompose market movements data **D** into principal components **P** and corresponding uncorrelated market dynamics **U** with weights **S**
- Use few SVD components with largest singular values low rank approximation of market data
- Principal components P are eigen vectors of covariance matrix D^TD

Main Principal Components of Swap Rates



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$\mathbf{P}^{\mathrm{T}}(\mathbf{r} + \mathbf{H}\mathbf{x}) = 0$
$\mathbf{x} = (\mathbf{P}^{\mathrm{T}}\mathbf{H})^{-1}\mathbf{P}^{\mathrm{T}}\mathbf{r}$
$\mathbf{x} = \mathbf{R}^{\mathrm{T}}\mathbf{r}$
$\mathbf{R} = \mathbf{P}(\mathbf{H}^{\mathrm{T}}\mathbf{P})^{-1}$

Swap	1Y	2Y	5Y	10Y	30Y
IRS=1Y	1	0	0	0	0
IRS=2Y	0	1	0	0	0
IRS=3Y	0	0	0	0	0
IRS=4Y	0	0	0	0	0
IRS=5Y	0	0	1	0	0
IRS=6Y	0	0	0	0	0
IRS=7Y	0	0	0	0	0
IRS=8Y	0	0	0	0	0
IRS=9Y	0	0	0	0	0
IRS=10Y	0	0	0	1	0
IRS=12Y	0	0	0	0	0
IRS=15Y	0	0	0	0	0
IRS=20Y	0	0	0	0	0
IRS=25Y	0	0	0	0	0
IRS=30Y	0	0	0	0	1

- **P** PCA factors
- H risk of hedging portfolio (liquid swaps)
- **R** risk transform matrix

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Hedging matrix H

Hedging using PCA model

		PCA Matrix				
Swap	Raw Risk	1Y	2Y	5Y	10Y	30Y
IRS=1Y	(200,000)	1.00	0.00	0.00	0.00	0.00
IRS=2Y	1,330,000	0.00	1.00	0.00	0.00	0.00
IRS=3Y	(200,000)	-0.51	1.16	0.29	-0.02	-0.02
IRS=4Y	1,200,000	-0.32	0.60	0.70	-0.04	-0.01
IRS=5Y	(722,450)	0.00	0.00	1.00	0.00	0.00
IRS=6Y	(35,255)	0.02	-0.05	0.80	0.28	-0.04
IRS=7Y	(537,430)	-0.01	-0.03	0.54	0.54	-0.04
IRS=8Y	(3,850,000)	-0.01	0.02	0.33	0.73	-0.05
IRS=9Y	1,580,000	-0.01	0.01	0.15	0.88	-0.03
IRS=10Y	288,751	0.00	0.00	0.00	1.00	0.00
IRS=12Y	(401,350)	0.02	0.01	-0.08	0.95	0.11
IRS=15Y	50,000	0.01	0.00	-0.04	0.59	0.45
IRS=20Y	4,000,000	0.01	0.03	-0.08	0.44	0.62
IRS=25Y	(1,000,000)	0.00	0.02	-0.04	0.20	0.82
IRS=30Y	(1,500,000)	0.00	0.00	0.00	0.00	1.00
	PCA risk	(426,757)	1,892,770	(1,538,808)	(268,764)	293,261
	B/O charge	0.1	0.1	0.1	0.1	0.1
	Charge	42,676	189,277	153,881	26,876	29,326

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- "Formally" tuned to historical data
- Hedge coefficients are not stable, especially if historical widow is short to reflect recent regime
- Costly to re-hedge when PC factors change
- Instability is coming from noisy PCs corresponding to small singular values
- Over-fitting to historical data
- No assumptions used about shape of the yield curve

Risk matrix **R** is linear combinations **Y** of principal components **P** producing shifts of one hedging instrument at a time

R = PY $H^{T}R = I$ $R = P(H^{T}P)^{-1}$

 Can we build risk model **R** based on some reasonable assumption (such as smoothness of forward rates) rather than purely historical data?

• Assumption: Forward rates move smoothly

 $H^{T}R = I$ || LJR ||² \rightarrow min R ~ (HH^T + λ^{2} (LJ)^T LJ)⁻¹

• Where J – Jacobian matrix translating shifts of yield curve inputs to movements of forward rates, L – smoothness regularity matrix, λ - small regularization parameter

Swap	Raw Risk	1Y	2Y	5Y	10Y	30Y
IRS=1Y	(200,000)	1.00	0.00	0.00	0.00	0.00
IRS=2Y	1,330,000	0.00	1.00	0.00	0.00	0.00
IRS=3Y	(200,000)	-0.28	0.95	0.44	-0.12	0.01
IRS=4Y	1,200,000	-0.13	0.36	0.92	-0.17	0.01
IRS=5Y	(722,450)	0.00	0.00	1.00	0.00	0.00
IRS=6Y	(35,255)	0.04	-0.11	0.81	0.28	-0.02
IRS=7Y	(537,430)	0.04	-0.12	0.58	0.53	-0.04
IRS=8Y	(3,850,000)	0.03	-0.09	0.35	0.75	-0.04
IRS=9Y	1,580,000	0.02	-0.04	0.15	0.90	-0.03
IRS=10Y	288,751	0.00	0.00	0.00	1.00	0.00
IRS=12Y	(401,350)	-0.02	0.05	-0.17	1.05	0.09
IRS=15Y	50,000	-0.03	0.07	-0.24	0.93	0.27
IRS=20Y	4,000,000	-0.02	0.05	-0.19	0.59	0.56
IRS=25Y	(1,000,000)	-0.01	0.03	-0.09	0.26	0.81
IRS=30Y	(1,500,000)	0.00	0.00	0.00	0.00	1.00
TOTAL	2266	(487,769)	2,082,997	(1,752,958)	78,962	64,483

Pricing Model Diagram



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Heath-Jarrow-Morton (HJM) Model

• Evolution of forward rates

$$df_{t,s} = \mu_{t,s} dt + f_{t,s}^{\beta} V(t,s) \rho(t,s) \cdot dB_{t}^{Q}$$

- *f* forward rate
- µ- drift
- β model skew factor
- ρ correlation/factor structure
- V(t,s) parametric volatility surface (our main focus today)
 - dB_{t}^{ϱ} Brownian motion

Forward Rates Map



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Parametric Volatility Surface



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Volatility Surface Calibration Challenge

- High dimensionality (need to calibrate ~28k elements)
- No memory to store 28k X 28k matrix
- Relatively small number of calibration instruments (20-50)
- Under-determined problem
- Sensitivity areas of calibration instruments overlap significantly
- Ill posed inverse problem
- Unstable, noisy solution
- Need regularity constraints
- Has to be smooth to produce realistic prices for similar instruments

Formal Approach to Calibration

 Represent volatility surface as a linear combination of N basis functions

 $\mathbf{v} = \mathbf{v}_0 + \mathbf{B} \cdot \mathbf{x}$

- v vector containing elements of the volatility grid
- B matrix, columns corresponding to basis fuctions
- x vector of weights
- Make N equivalent to the number of calibration instruments M
- "Formally" unambiguous
- Make basis functions piecewise constant matching sensitivity of calibration instruments, 0 otherwise

Compute sensitivities (Jacobian matrix)

• Use pricing model to compute sensitivities of prices of calibration instruments to perturbations of volatility surface

$$J_{ij} = \frac{\partial q_i}{\partial x_j}$$
$$\mathbf{q} = \mathbf{J} \cdot \mathbf{x}$$
$$\mathbf{q} = \ln \frac{\mathbf{q}_{mdl}}{\mathbf{q}_0}, \mathbf{q}_{in} = \ln \frac{\mathbf{q}_{market}}{\mathbf{q}_0}$$

Where **J** – Jacobian matrix

 \mathbf{q}_{mdl} , \mathbf{q}_{market} , \mathbf{q}_0 – model, market, and base price \mathbf{x} – vector of basis functions coefficients

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- J is square and invertible, as basis functions are selected reasonably
- Iteratively solve for basis function coefficients ${\boldsymbol x}$

$$\mathbf{q}_{in} = \mathbf{J} \cdot \mathbf{x}$$

 $\mathbf{x} = \mathbf{J}^{-1} \mathbf{q}_{in}$

 Quickly converges, as (typically) price is ~proportional to volatility for at-the-money calibration instruments

"Formal" solution

• Exact, but ... meaningless



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Attempt to improve solution



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Calibration problem: Ill-posed



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Key Improvements of Calibration

- Use ill-posedness to our advantage:
 - Allow some tolerance to calibration accuracy of input instruments
 - Significant improvements in the (smoothness of) output surface may not cost much in terms of accuracy of calibration
 - Calibration instruments have different liquidity, and bid-offer spread. So we can use weights to decrease tolerance for important instruments
- Basis functions:
 - Absolutely need basis functions to reduce dimensionality of the inverse problem
 - Need many (M > N instruments) basis functions, as we do not know in advance which shapes will work
 - 2-dimensional B-Splines

B-Spline representation



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Building B-splines using Cox-de Boor recursion formula



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2-d B-spline functions



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$$\| \mathbf{W} \cdot (\mathbf{q} - \mathbf{q}_{in}) \|^2 \rightarrow \min$$
$$\| \mathbf{L}_1 \cdot (\mathbf{v} - \mathbf{v}_0) \|^2 \rightarrow \min$$
$$\| \mathbf{L}_2 \cdot \mathbf{v} \|^2 \rightarrow \min$$

Where W – diagonal matrix of weights

- L_1 regularization matrix for change
- L_2 regularization matrix for result

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Surface Gradient Penalty

Example of regularization matrix L_1

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$$\mathbf{x} = \arg\min\{\| \mathbf{W}(\mathbf{J}\mathbf{x} - \mathbf{q}_{in})\|^{2} + \| \lambda_{1}\mathbf{L}_{1}\mathbf{B}\mathbf{x}\|^{2} + \| \lambda_{2}\mathbf{L}_{2}\mathbf{x}\|^{2} \}$$
$$\mathbf{x} = \mathbf{A} \cdot \mathbf{J}^{T}\mathbf{W}^{2}\mathbf{q}_{in}, where$$
$$\mathbf{A} = (\mathbf{J}^{T}\mathbf{W}^{2}\mathbf{J} + \lambda_{1}^{2}(\mathbf{L}_{1}\mathbf{B})^{T}\mathbf{L}_{1}\mathbf{B} + \lambda_{2}^{2}\mathbf{L}_{2}^{T}\mathbf{L}_{2})^{-1}$$

Where L_2 – Tikhonov regularization matrix

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Calibration Result



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Calibration Inverse Problem

$$\mathbf{y} = \mathbf{A}\mathbf{x} + \varepsilon$$
$$\mathbf{x} = (\mathbf{A}^{\mathrm{T}}\mathbf{A})^{-1}\mathbf{A}^{\mathrm{T}}\mathbf{y}$$
$$\mathbf{A} = \mathbf{U}\mathbf{S}\mathbf{V}^{\mathrm{T}} = \sum_{i} s_{i}\mathbf{u}_{i}\mathbf{v}_{i}^{\mathrm{T}}$$
$$\mathbf{U}^{\mathrm{T}}\mathbf{U} = \mathbf{V}^{\mathrm{T}}\mathbf{V} = \mathbf{I}$$
$$\mathbf{x} = \mathbf{V}\mathbf{S}^{-1}\mathbf{U}^{\mathrm{T}}\mathbf{y} = \sum_{i} \frac{\mathbf{u}_{i}^{\mathrm{T}}\mathbf{y}}{s_{i}}\mathbf{v}$$

- y market inputs, x model parameters
- Singular Value Decomposition of (forward) model matrix **A**
- s_i singular values
- Result: Rotation \rightarrow Scaling $1/s_i \rightarrow$ Rotation

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$$\mathbf{x} = \sum_{i} \frac{\mathbf{u}_{i}^{\mathrm{T}} \mathbf{\varepsilon}}{S_{i}} \mathbf{v}_{i}$$

- Input noise εmay be magnified by small singular values s_i
- Condition number $max(s_i)/min(s_i)$ as indicator of ill-posedness
- Small variation in input results in large change in the solution

$$\mathbf{x} = \sum_{i} \frac{\mathbf{u}_{i}^{\mathrm{T}} \mathbf{y}}{S_{i}} \mathbf{v}_{i}$$
$$\mathbf{p} = \mathbf{B} \mathbf{x}$$
$$\mathbf{p}_{y} = \mathbf{A} \mathbf{x} = \mathbf{A} (\mathbf{A}^{\mathrm{T}} \mathbf{A})^{-1} \mathbf{A}^{\mathrm{T}} \mathbf{y} = \mathbf{y}$$

- We compute ${f x}$ to calculate price of the portfolio ${f p}$
- If all singular values are non-zero, we "formally" re-price inputs, as
 A^T(A^TA)⁻¹A=I identity. So, the model appears to be accurate.
- However, pricing of actual portfolio **p** with model **B** may be unstable

$$\mathbf{x} = \arg\min\{||\mathbf{A}\mathbf{x} - \mathbf{y}||^{2} + ||\lambda\mathbf{x}||^{2}\}$$
$$\mathbf{x} = (\mathbf{A}^{\mathrm{T}}\mathbf{A} + \lambda^{2}\mathbf{I})^{-1}\mathbf{A}^{\mathrm{T}}\mathbf{y}$$
$$\mathbf{x} = \mathbf{V}\frac{\mathbf{S}}{\mathbf{S}^{2} + \lambda^{2}\mathbf{I}}\mathbf{U}^{\mathrm{T}}\mathbf{y} = \sum_{i} w_{i} \cdot \frac{\mathbf{u}_{i}^{\mathrm{T}}\mathbf{y}}{s_{i}}\mathbf{v}_{i}$$
$$w_{i} = \frac{s_{i}^{2}}{s_{i}^{2} + \lambda^{2}}$$

- Apply penalty to amplitude of model parameters ${\boldsymbol x}$
- Re-pricing matrix A^T(A^TA+λ²I)⁻¹A≠I is no longer 100% accurate, however
- More stable model vector **x**, and pricing of actual portfolio

Truncated SVD (TSVD)

$$\mathbf{x} = \sum_{i} w_{i} \cdot \frac{\mathbf{u}_{i}^{\mathrm{T}} \mathbf{y}}{S_{i}} \mathbf{v}_{i}$$
$$w_{i} = 1, i \leq N$$
$$w_{i} = 0, i > N$$

- Truncation of effective rank of the model matrix ${\boldsymbol{\mathsf{A}}}$
- Similarity with PCA (principal component) approach
- Truncated is "null" space of the model: parameter modes, which do not affect calibration accuracy
- The problem is when "null" space of the model has noticeable impact on portfolio pricing

- Improved stability
- Regularization is essential for ill-conditioned problems
- More realistic solution at the expense of fitting input data
- May cause bias to the solution
- Bias can be minimized by proper selection of the penalty constraints

Useful Links

- HJM model: http://en.wikipedia.org/wiki/HJM_model
- Yield Curve: http://en.wikipedia.org/wiki/Yield_curve
- Inverse problems: http://en.wikipedia.org/wiki/Inverse_problem
- Tikhonov Regularization: http://en.wikipedia.org/wiki/Tikhonov_regularization
- Singular Value decomposition:

http://en.wikipedia.org/wiki/Singular_value_decomposition

- PCA: http://en.wikipedia.org/wiki/Principal_component_analysis
- B-Splines: http://en.wikipedia.org/wiki/B-spline

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