## Manipulating Continuous Random Variables Class 5, 18.05, Spring 2014 Jeremy Orloff and Jonathan Bloom

## 1 Learning Goals

1. Be able to find the pdf and cdf of a random variable defined in terms of a random variable with known pdf and cdf.

## 2 Transformations of Random Variables

If Y = aX + b then the properties of expectation and variance tell us that E(Y) = aE(X) + band  $Var(Y) = a^2Var(X)$ . But what is the distribution function of Y? If Y is continuous, what is its pdf?

Often, when looking at transforms of discrete random variables we work with tables. For continuous random variables transforming the pdf is just change of variables ('*u*-substitution') from calculus. Transforming the cdf makes direct use of the definition of the cdf.

Let's remind ourselves of the basics:

1. The cdf of X is  $F_X(x) = P(X \le x)$ .

2. The pdf of X is related to  $F_X$  by  $f_X(x) = F'_X(x)$ .

**Example 1.** Let  $X \sim U(0,2)$ , so  $f_X(x) = 1/2$  and  $F_X(x) = x/2$  on [0,2]. What is the range, pdf and cdf of  $Y = X^2$ ?

**answer:** The range is easy: [0, 4].

To find the cdf we work systematically from the definition.

$$F_Y(y) = P(Y \le y) = P(X^2 \le y) = P(X \le \sqrt{y}) = F_X(\sqrt{y}) = \sqrt{y}/2.$$

To find the pdf we can just differentiate the cdf

$$f_Y(y) = \frac{d}{dy}F_Y(y) = \boxed{\frac{1}{4\sqrt{y}}}.$$

An alternative way to find the pdf directly is by change of variables. The trick here is to remember that it is  $f_X(x)dx$  which gives probability  $(f_X(x))$  by itself is probability density). Here is how the calculation goes in this example.

$$y = x^2 \Rightarrow dy = 2x \, dx \Rightarrow dx = \frac{dy}{2\sqrt{y}}$$
  
 $f_X(x) \, dx = \frac{dx}{2} = \frac{dy}{4\sqrt{y}} = f_Y(y) \, dy$ 

Therefore  $f_Y(y) = \frac{dy}{4\sqrt{y}}$ 

**Example 2.** Let  $X \sim \exp(\lambda)$ , so  $f_X(x) = \lambda e^{-\lambda x}$  on  $[0, \infty]$ . What is the density of  $Y = X^2$ ? **answer:** Let's do this using the change of variables.

$$y = x^2 \Rightarrow dy = 2x \, dx \Rightarrow dx = \frac{dy}{2\sqrt{y}}$$
  
 $f_X(x) \, dx = \lambda e^{-\lambda x} \, dx = \lambda e^{-\lambda \sqrt{y}} \frac{dy}{2\sqrt{y}} = f_Y(y) \, dy$ 

Therefore  $f_Y(y) = \frac{\lambda}{2\sqrt{y}} e^{-\lambda\sqrt{y}}$ .

**Example 3.** Assume  $X \sim N(5, 3^2)$ . Show that  $Z = \frac{X-5}{3}$  is standard normal, i.e.,  $Z \sim N(0, 1)$ .

**answer:** Again using the change of variables and the formula for  $f_X(x)$  we have

$$z = \frac{x-5}{3} \Rightarrow dz = \frac{dx}{3} \Rightarrow dx = 3 dz$$

$$f_X(x) \, dx = \frac{1}{3\sqrt{2\pi}} \mathrm{e}^{-(x-5)^2/(2\cdot 3^2)} \, dx = \frac{1}{3\sqrt{2\pi}} \mathrm{e}^{-z^2/2} \, 3 \, dz = \frac{1}{\sqrt{2\pi}} \mathrm{e}^{-z^2/2} \, dz = f_Z(z) \, dz$$

Therefore  $f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$ . Since this is exactly the density for N(0,1) we have shown that Z is standard normal.

This example shows an important general property of normal random variables which we give in the next example.

**Example 4.** Assume  $X \sim N(\mu, \sigma^2)$ . Show that  $Z = \frac{X - \mu}{\sigma}$  is standard normal, i.e.,  $Z \sim N(0, 1)$ .

**answer:** This is exactly the same computation as the previous example with  $\mu$  replacing 5 and  $\sigma$  replacing 3. We show the computation without comment.

$$z = \frac{x - \mu}{\sigma} \Rightarrow dz = \frac{dx}{\sigma} \Rightarrow dx = \sigma dz$$

$$f_X(x) \, dx = \frac{1}{\sigma\sqrt{2\pi}} \mathrm{e}^{-(x-\mu)^2/(2\cdot\sigma^2)} \, dx = \frac{1}{\sigma\sqrt{2\pi}} \mathrm{e}^{-z^2/2} \, \sigma \, dz = \frac{1}{\sqrt{2\pi}} \mathrm{e}^{-z^2/2} \, dz = f_Z(z) \, dz$$

Therefore  $f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$ . This shows Z is standard normal.

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