### 18.05 Problem Set 2, Spring 2014 Solutions

Problem 1. (10 pts.)
(a) Listing the gender of older child first, our sample space is $\{B B, B G, G B, G G\}$. The event "the older child is a girl" is $\{G B, G G\}$ and the event "both children are girls" is $\{G G\}$. Thus the probability that both children are girls given the older child is a girl is $\frac{1}{2}$.
(b) The event "at least one child is a boy" is $\{B B, B G, G B\}$ so the probability that both children are boys is $\frac{1}{3}$.

Problem 2. (10 pts.) The defense will try to make the case that this is likely to be a case of random mis-identification. So we look for the probability a random taxi the witness sees as blue is actually blue.
This is a question of 'inverting' conditional probability. We know

$$
P(\text { the witness sees blue|the car is blue) }
$$

but we'd like to know
$P$ (the car is blue|the witness sees blue).
Our first job is to translate this to symbols.
Let $W_{b}=$ 'witness sees a blue taxi' and let $W_{g}=$ 'witness sees a green car'. Further, let $T_{b}$ $=$ 'taxi is blue' and let $T_{g}=$ 'taxi is green'. With this notation we want to find $P\left(T_{b} \mid W_{b}\right)$. We will compute this using Bayes' formula

$$
P\left(T_{b} \mid W_{b}\right)=\frac{P\left(W_{b} \mid T_{b}\right) \cdot P\left(T_{b}\right)}{P\left(W_{b}\right)}
$$

All the pieces are represented in the following diagram.

Random taxi
Witness sees


We can determine each factor in the right side of Bayes' formula:
We are given $P\left(T_{b}\right)=.01$ (and $\left.P\left(T_{g}\right)=.99\right)$.
We are given, $P\left(W_{b} \mid T_{b}\right)=.99$ and $P\left(W_{b} \mid T_{g}\right)=.02$.
We compute $P\left(W_{b}\right)$ using the law of total probability:

$$
P\left(W_{b}\right)=P\left(W_{b} \mid T_{b}\right) P\left(T_{b}\right)+P\left(W_{b} \mid T_{g}\right) P\left(T_{g}\right)=.99 \times .01+.02 \times .99=.99 \times .03
$$

Putting all this in Bayes' formula we get

$$
P\left(T_{b} \mid W_{b}\right)=\frac{.99 \times .01}{.99 \times .03}=\frac{1}{3}
$$

Ladies and gentlemen of the jury. The prosecutor tells you that the witness is nearly flawless in his ability to distinguish whether a taxi is green or blue. He claims that this
implies that beyond a reasonable doubt the taxi involved in the hit and run was blue. However probability theory shows without any doubt that the probability a random taxi seen by the witness as blue is actually blue is only $1 / 3$. This is considerably more than a reasonable doubt. In fact it is more probable than not that the taxi involved in the accident was green. If the probability doesn't fit you must acquit!

Problem 3. ( 10 pts .) The following tree shows all the possible values for $X$ (note, below each edge in the tree is the conditional probability that edge).


At this point we have enough information to compute expectation.

$$
E(X)=2\left(\frac{3}{14}\right)+3\left(\frac{2}{7}\right)+4\left(\frac{1}{7}\right)+5\left(\frac{1}{7}\right)+5\left(\frac{1}{7}\right)+6\left(\frac{1}{14}\right)=\frac{26}{7} \approx 3.7143
$$

We can also give the probability distribution of $X$

| X | 2 | 3 | 4 | 5 | 6 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{p}(\mathrm{X})$ | $3 / 14$ | $2 / 7$ | $1 / 7$ | $2 / 7$ | $1 / 14$ |

Problem 4. (10 pts.) (a)

| S | 4 | 6 | 8 |
| :---: | :---: | :---: | :---: |
| $\mathrm{p}(\mathrm{X})$ | $1 / 4$ | $1 / 4$ | $1 / 2$ |

(b) Bayes' rule says

$$
P(S=k \mid R=3)=\frac{P(R=3 \mid S=k) P(S=k)}{P(R=3)} .
$$

We summarize what we know in a tree. In the tree the notation $S_{4}$ means the 4 -sided die ( $S=4$ ), likewise $R_{3}$ means a 3 was rolled $(R=3)$. Because we only care about the case $R=3$ the tree does not include other possible rolls.

Chosen die
Roll result


We have the following probabilities (you should identify them in the tree):

$$
P(R=3 \mid S=4)=1 / 4, \quad P(R=3 \mid S=6)=1 / 6, \quad P(R=3 \mid S=8)=1 / 8 .
$$

The law of total probability gives (again, see how the tree tells us this):

$$
\begin{aligned}
P(R=3) & =P(R=3 \mid S=4) P(S=4)+P(R=3 \mid S=6) P(S=6)+P(R=3 \mid S=8) P(S=8) \\
& =\frac{1}{4} \cdot \frac{1}{4}+\frac{1}{6} \cdot \frac{1}{4}+\frac{1}{8} \cdot \frac{1}{2} \\
& =\frac{1}{6} .
\end{aligned}
$$

Hence

$$
\begin{aligned}
& P(S=4 \mid R=3)=\frac{1 / 16}{1 / 6}=\frac{3}{8} \\
& P(S=6 \mid R=3)=\frac{1 / 24}{1 / 6}=\frac{1}{4} \\
& P(S=8 \mid R=3)=\frac{1 / 16}{1 / 6}=\frac{3}{8} .
\end{aligned}
$$

(c) In a similar vein, we have

$$
P(R=6 \mid S=4)=0, \quad P(R=6 \mid S=6)=1 / 6, \quad P(R=6 \mid S=8)=1 / 8 .
$$

and

$$
P(R=6)=0 \cdot \frac{1}{4}+\frac{1}{6} \cdot \frac{1}{4}+\frac{1}{8} \cdot \frac{1}{2}=\frac{5}{48} .
$$

So,

$$
P(S=4 \mid R=6)=0, \quad P(S=6 \mid R=6)=\frac{1 / 24}{5 / 48}=\frac{2}{5}, \quad P(S=8 \mid R=6)=\frac{1 / 16}{5 / 48}=\frac{3}{5} .
$$

The eight-sided die is more likely. Note, the denominator is the same in each probability, i.e. the total probability $P(R=6)$, so all we had to check was the numerator.
(d) The only way to get $R=7$ is if we picked an octahedral die.

Problem 5. ( 10 pts .) Label the seats 1 to $n$ going clockwise around the table. Let $X_{i}$ be the Bernoulli random variable with value 1 if the person in seat $i$ is shorter than his or her neighbors. Then $X=\sum_{i=1}^{n} X_{i}$ represents the total number of people who are shorter than both of their neighbors, and

$$
E(X)=E\left(\sum_{i=1}^{n} X_{i}\right)=\sum_{i=1}^{n} E\left(X_{i}\right)
$$

by linearity of expected value. Recall that this property of expected values holds even when the $X_{i}$ are dependent, as is the case here!

Among 3 random people the probability that the middle one is the shortest is $1 / 3$. Therefore $X_{i} \sim \operatorname{Bernoulli}(1 / 3)$, which implies $E\left(X_{i}\right)=1 / 3$. Therefore the expected number of people shorter than both their neighbors is

$$
E(X)=\sum_{i=1}^{n} E\left(X_{i}\right)=\frac{n}{3} .
$$

## Alternate (and more complicated) solution

Suppose we number the $n$ people based on their height, from shortest to tallest (so person 1 is the shortest and person $n$ is the tallest). For person $k$, there are $n-k$ people who are taller. In order for person $k$ 's neighbors to be taller than person $k$, we can pick two neighbors from these $n-k$ in $\binom{n-k}{2}$ ways. Moreover, there are $\binom{n-1}{2}$ total ways to pick the neighbor of person $k$. Thus, the probability that person $k$ is seated between two people taller than him/her is

$$
p_{k}=\frac{\binom{n-k}{2}}{\binom{n-1}{2}} .
$$

This formula is valid for $k=1,2, \ldots, n-2$ and for $k=n-1$ and $k=n$, we let $p_{k}=0$ (since there is no way for the tallest person or the second tallest person to be sitting next to two people taller than themselves).

We can define $n$ Bernoulli random variables, $X_{1}, \ldots, X_{n}$, as follows:

$$
X_{k}= \begin{cases}1 & \text { if person } k \text { 's neighbors are taller than person } \mathrm{k} \\ 0 & \text { otherwise }\end{cases}
$$

Thus, $E\left(X_{k}\right)=p_{k}$ for $k=1,2, \ldots, n$. The total number of people who are seated next to people taller than them is $X=X_{1}+\cdots+X_{n-2}$, (again we ignore $X_{n-1}$ and $X_{n}$ since they have to be 0 ). So, we get, by linearity of expectation

$$
E(X)=\sum_{k=1}^{n-2} \frac{\binom{n-k}{2}}{\binom{n-1}{2}}
$$

One can show (after a bit of algebra) that this sum is equal to $n / 3$.
Had we ordered the people from tallest to shortest, then we would have

$$
E(X)=\sum_{k=3}^{n} \frac{\binom{k-1}{2}}{\binom{n-1}{2}}
$$

Problem 6. (10 pts.) (a) Any sequence of 500 's and 1 's is valid. However, most people do not put in any long runs that parts (c) and (d) will show happen frequently.
(c) Here is my code with comments
nflips = 50
ntrials $=10000$
total $=0$ \# We'll keep a running total of all the trials' longest runs

```
for (j in 1:ntrials)
{
    # One trial consists of 50 flips
    trial = rbinom(nflips, 1, .5) # binomial(1,.5) = bernoulli(.5)
    # rle() finds the lengths of all the runs in trials. We add the max to total
    total = total + max(rle(trial)$lengths)
}
# The average maximum run is the total/ntrials
aveMax = total/ntrials
print(aveMax)
My run of this code produced aveMax \(=5.9645\)
(d) Instead of keeping a total we keep a count of the number of trials with a run of 8 or more
```

```
nflips = 50
```

nflips = 50
ntrials = 10000

# We'll keep a running count of all the trials with a run of 8 or more

count = 0
for (j in 1:ntrials)
{
trial = rbinom(nflips, 1, .5) \# binomial(1,.5) = bernoulli(.5)
count = count + (max(rle(trial)\$lengths) >= 8)
}

# The probability of a run of 8 or more is count/ntrials

prob8 = count/ntrials
print(prob8)

```

My run of this code produced prob8 \(=0.1618\)

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\subsection*{18.05 Introduction to Probability and Statistics}

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